

Supplementary material for “Entropy-regularized mean-variance portfolio optimization with jumps”

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Abstract

This document contains supplementary materials for the main article [1]. All notations used here are in accordance with that in [1].

1 Integrability for solutions of SDEs with jumps

Although the following fact can be easily extended to a multidimensional setting, however, we formulate it in the one-dimensional case for the sake of simplicity.

Lemma 1.1. *Let $\xi = (\xi_t)_{t \in [0, T]}$ be càdlàg and adapted with $\|\xi\|_{\mathcal{S}_2([0, T])}^2 := \mathbb{E}[\sup_{0 \leq t \leq T} \xi_t^2] < \infty$. Assume that $dZ_t = \phi_t dt + dK_t$, where $K = (K_t)_{t \in [0, T]}$ is a càdlàg $\mathbf{L}_2(\mathbb{P})$ -martingale satisfying $d\langle K, K \rangle_t = \eta_t^2 dt$, where η and ϕ are progressively measurable with $\sup_{0 < t < T} \eta_t^2 + \int_0^T \phi_t^2 dt \leq C$ a.s. for some (non-random) constant $C > 0$. Then, for a Lipschitz function $\sigma: \mathbb{R} \rightarrow \mathbb{R}$, the SDE*

$$X_t = \xi_t + \int_0^t \sigma(X_{u-}) dZ_u, \quad X_0 = \xi_0 = x_0 \in \mathbb{R}, \quad (1.1)$$

has a unique càdlàg strong solution $X = (X_t)_{t \in [0, T]}$ satisfying $\mathbb{E}[\sup_{0 \leq t \leq T} X_t^2] \leq C' < \infty$ for some constant $C' = C'(\|\xi\|_{\mathcal{S}_2([0, T])}, T, \sigma, C) > 0$.

Proof. Due to [4, Theorem V.3.7], the SDE (1.1) has a unique càdlàg and adapted solution X . For $n \geq 1$ we define $\tau_n := \inf\{t > 0 : |X_t| \geq n\} \wedge T$. Then τ_n is a stopping time with $|X_{(t \wedge \tau_n)-}| \leq n$ for $t \in [0, T]$. It is known that, see, e.g., [4, Theorem II.5.12], a.s.,

$$X_{t \wedge \tau_n} = \xi_{t \wedge \tau_n} + \int_0^{t \wedge \tau_n} \sigma(X_{u-}) dZ_u = \xi_{t \wedge \tau_n} + \int_0^t \mathbf{1}_{(0, \tau_n]}(u) \sigma(X_{u-}) dK_u + \int_0^t \mathbf{1}_{(0, \tau_n]}(u) \sigma(X_{u-}) \phi_u du$$

so that the triangle inequality, Itô’s isometry, and Hölder’s inequality yield

$$\begin{aligned} \frac{1}{3} \mathbb{E}[X_{t \wedge \tau_n}^2] &\leq \mathbb{E}[\xi_{t \wedge \tau_n}^2] + \mathbb{E} \left[\int_0^t \mathbf{1}_{(0, \tau_n]}(u) \sigma(X_{u-})^2 \eta_u^2 du \right] + \mathbb{E} \left[\left| \int_0^t \mathbf{1}_{(0, \tau_n]}(u) \sigma(X_{u-}) \phi_u du \right|^2 \right] \\ &\leq \|\xi\|_{\mathcal{S}_2([0, T])}^2 + C \mathbb{E} \left[\int_0^t \mathbf{1}_{(0, \tau_n]}(u) \sigma(X_{u-})^2 du \right] \\ &\leq \|\xi\|_{\mathcal{S}_2([0, T])}^2 + 2CT\sigma(0)^2 + 2C|\sigma|_{\text{Lip}}^2 \mathbb{E} \left[\int_0^t \mathbf{1}_{(0, \tau_n]}(u) X_{u-}^2 du \right] \\ &= \alpha + \beta \mathbb{E} \left[\int_0^t \mathbf{1}_{(0, \tau_n]}(u) X_{u-}^2 du \right] \end{aligned}$$

for $\alpha := \|\xi\|_{\mathcal{S}_2([0, T])}^2 + 2CT\sigma(0)^2$, $\beta := 2C|\sigma|_{\text{Lip}}^2$, and $|\sigma|_{\text{Lip}} := \sup_{x \neq y} \frac{|\sigma(x) - \sigma(y)|}{|x - y|}$. Since

$$\mathbf{1}_{(0, \tau_n]}(u) X_{u-}^2 \leq n^2, \quad u \in [0, T],$$

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it implies that

$$\mathbb{E}[X_{t \wedge \tau_n}^2] \leq 3\alpha + 3\beta T n^2, \quad \forall t \in [0, T].$$

Moreover, as X has càdlàg paths, we get for all $t \in [0, T]$ that

$$\frac{1}{3}\mathbb{E}[X_{t \wedge \tau_n}^2] \leq \alpha + \beta \mathbb{E}\left[\int_0^t \mathbf{1}_{(0, \tau_n)}(u) X_u^2 du\right] \leq \alpha + \beta \mathbb{E}\left[\int_0^t X_{u \wedge \tau_n}^2 du\right] = \alpha + \beta \int_0^t \mathbb{E}[X_{u \wedge \tau_n}^2] du.$$

Applying Gronwall's lemma yields $\mathbb{E}[X_{t \wedge \tau_n}^2] \leq 3\alpha e^{3\beta T}$ for all $t \in [0, T]$, $n \geq 1$. Since $(\tau_n)_{n \geq 1}$ is eventually constant T a.s., sending $n \rightarrow \infty$ and using Fatou's lemma we obtain

$$\mathbb{E}[X_t^2] \leq \liminf_{n \rightarrow \infty} \mathbb{E}[X_{t \wedge \tau_n}^2] \leq 3\alpha e^{3\beta T}, \quad \forall t \in [0, T].$$

As a consequence, $\int_0^\cdot \sigma(X_{t-}) dK_t$ is an $\mathbf{L}_2(\mathbb{P})$ -martingale. Therefore, applying Doob's maximal inequality for the martingale part we get

$$\mathbb{E}\left[\sup_{0 \leq t \leq T} X_t^2\right] \leq \|\xi\|_{\mathcal{S}_2([0, T])}^2 + 4\mathbb{E}\left[\int_0^T \sigma(X_{t-})^2 \eta_t^2 dt\right] + \mathbb{E}\left[\int_0^T \sigma(X_{t-})^2 dt \int_0^T \phi_t^2 dt\right] < \infty,$$

which completes the proof. \square

2 Explicit expression for the optimal wealth and Lagrange multiplier

We give in this part a closed-form representation of the optimal wealth X^* and the respective Lagrange multiplier \hat{w} when the condition “ $\Delta Z \neq 1$ on $[0, T]$ ” in [1, Proposition 4.12] fails to hold. Let us impose the assumptions of [1, Theorem 4.10]. For Z given in that theorem, we write

$$-(X_{s-}^* - \hat{w})dZ_s = (X_{s-}^* - \hat{w})d(-Z_s)$$

and follow [4, Exercise V.27] to define the sequence of stopping times $\{\tau_n\}_{n \geq 1}$ by setting

$$\tau_0 := 0, \quad \tau_n := \inf\{\tau_{n-1} < s \leq T : 1 - \Delta Z_s = 0\}, \quad n \geq 1.$$

Note that τ_n is non-decreasing and tends to ∞ a.s. as $n \rightarrow \infty$. Then the solution X^* of [1, Eq. (4.23)] is

$$X_r^* = \hat{w} + \sum_{n=1}^{\infty} X_r^{*,n-1} \mathbf{1}_{[\tau_{n-1}, \tau_n) \cap [0, T]}(r), \quad r \in [0, T], \quad (2.1)$$

where we conventionally set $[\infty, \infty) := \emptyset$. In (2.1), $X_r^{*,n-1} = (X_r^{*,n-1})_{r \in [0, T]}$ is given by

$$X_r^{*,n-1} := \left[(x_0 - \hat{w}) \mathbf{1}_{\{n=1\}} + \sqrt{\frac{\lambda}{2}} \left(\Delta M_{\tau_{n-1}} + \int_{\tau_{n-1}}^r \frac{dM_s}{U_{s-}^{n-1}} + \int_{\tau_{n-1}}^r \frac{d[M, Z]_s}{U_{s-}^{n-1}} \right) \right] U_r^{n-1} \mathbf{1}_{[\tau_{n-1}, \tau_n) \cap [0, T]}(r).$$

The process $U^{n-1} = (U_r^{n-1})_{r \in [0, T]}$ is defined by

$$U_r^{n-1} := \begin{cases} 1 & \text{if } r \leq \tau_{n-1} \\ e^{-Z_r + Z_{\tau_{n-1}} - \frac{1}{2} \int_{\tau_{n-1}}^r d[Z, Z]_s^c} \prod_{\tau_{n-1} < s \leq r} (1 - \Delta Z_s) e^{\Delta Z_s} & \text{if } r > \tau_{n-1} \end{cases} \\ = \mathcal{E}(-Z + Z^{\tau_{n-1}})_r,$$

where $\mathcal{E}(-Z + Z^{\tau_{n-1}}) = (\mathcal{E}(-Z + Z^{\tau_{n-1}})_r)_{r \in [0, T]}$ denotes the Doléans–Dade exponential of $-Z + Z^{\tau_{n-1}}$, see [4, Section II.8], and where $Z^{\tau_{n-1}}$ is the process Z stopped at τ_{n-1} , i.e. $Z_t^{\tau_{n-1}} := Z_{t \wedge \tau_{n-1}}$.

We now calculate the Lagrange multiplier \hat{w} using the constraint $\mathbb{E}[X_T^*] = \hat{z}$. One first has

$$d\mathcal{E}(-Z + Z^{\tau_{n-1}})_s = \mathcal{E}(-Z + Z^{\tau_{n-1}})_{s-} d(-Z_s + Z_s^{\tau_{n-1}}), \quad \mathcal{E}(-Z + Z^{\tau_{n-1}})_0 = 1,$$

and the *conditional* quadratic variation¹ of $-Z + Z^{\tau_{n-1}}$, which is computed by

$$\begin{aligned} & \langle -Z + Z^{\tau_{n-1}}, -Z + Z^{\tau_{n-1}} \rangle \\ &= \left\langle \int_0^{\cdot} \mathbb{1}_{(\tau_{n-1}, T]}(s) (\mathcal{M}_\alpha^\top \mathcal{S}_\alpha^{-1})(s, Y_{s-}) dY_s, \int_0^{\cdot} \mathbb{1}_{(\tau_{n-1}, T]}(s) (\mathcal{M}_\alpha^\top \mathcal{S}_\alpha^{-1})(s, Y_{s-}) dY_s \right\rangle \\ &= \int_0^{\cdot} \mathbb{1}_{(\tau_{n-1}, T]}(s) (\mathcal{M}_\alpha^\top \mathcal{S}_\alpha^{-1} \Sigma \mathcal{S}_\alpha^{-1} \mathcal{M}_\alpha)(s, Y_{s-}) ds, \end{aligned}$$

has uniformly bounded integrand over $(n, s) \in \mathbb{N} \times (0, T)$ a.s. by [1, Eq. (4.20)]. Then applying Lemma 1.1 yields $\sup_{n \geq 1} \mathbb{E}[|U_T^{n-1}|^2] = \sup_{n \geq 1} \mathbb{E}[|\mathcal{E}(-Z + Z^{\tau_{n-1}})_T|^2] < \infty$. In particular, for $n = 1$ we can define

$$d_{(2.2)} := \mathbb{E}[\mathcal{E}(-Z)_T \mathbb{1}_{[0, \tau_1] \cap [0, T]}(T)] = \mathbb{E}[\mathcal{E}(-Z)_T \mathbb{1}_{\{\tau_1 = \infty\}}] \in \mathbb{R}. \quad (2.2)$$

Moreover, using Lemma 1.1 again we assert that X^* is a square integrable process which together with (2.1) and (2.2) then imply that

$$d_{(2.3)} := \mathbb{E} \left[\sum_{n=1}^{\infty} \left(\Delta M_{\tau_{n-1}} + \int_{\tau_{n-1}}^T \frac{dM_s}{U_{s-}^{n-1}} + \int_{\tau_{n-1}}^T \frac{d[M, Z]_s}{U_{s-}^{n-1}} \right) U_T^{n-1} \mathbb{1}_{[\tau_{n-1}, \tau_n) \cap [0, T]}(T) \right] \quad (2.3)$$

finitely exists. Now we let $r = T$ and take the expectation both sides of (2.1) to get

$$\hat{z} = \hat{w} + d_{(2.2)}(x_0 - \hat{w}) + \sqrt{\frac{\lambda}{2}} d_{(2.3)},$$

If $d_{(2.2)} \neq 1$, then the Lagrange multiplier \hat{w} is calculated by

$$\hat{w} = \frac{\hat{z} - \sqrt{\frac{\lambda}{2}} d_{(2.3)} - d_{(2.2)} x_0}{1 - d_{(2.2)}}.$$

3 Some auxiliary results for Lévy processes

All Lévy processes below are considered with the canonical truncation function $h(x) = x \mathbb{1}_{\{\|x\| \leq 1\}}$.

Lemma 3.1. *L is a D -dimensional Lévy process if and only if $u^\top L$ is a 1-dimensional Lévy process for all $u \in \mathbb{R}^D$. Moreover, L has characteristic (b, A, ν) if and only if $u^\top L$ has characteristic $(b_u, u^\top A u, \nu \circ \{y \mapsto u^\top y\}^{-1})$ where $b_u := u^\top b - \int_{u^\top y \neq 0} u^\top y (\mathbb{1}_{\{\|y\| \leq 1\}} - \mathbb{1}_{\{|u^\top y| \leq 1\}}) \nu(dy)$ for all $u \in \mathbb{R}^D$.*

Proof. It is obvious that L has càdlàg paths a.s. if and only if $u^\top L$ has càdlàg paths a.s. for all $u \in \mathbb{R}^D$. We now verify the equivalence regarding distributional properties. Let $\mathcal{F}_t^L := \sigma\{L_s : s \leq t\}$. Assume that L is a D -dimensional Lévy process with characteristic (b, A, ν) . Then it follows from [3, Theorem 3.1] that, for any $s \leq t$ and $x \in \mathbb{R}$, a.s.,

$$\mathbb{E}[e^{ixu^\top(L_t - L_s)} | \mathcal{F}_s^L] = e^{-(t-s)\kappa(xu)}.$$

By a change of variables we have

$$\begin{aligned} \kappa_u(x) &:= \kappa(xu) = -ixu^\top b + \frac{x^2 u^\top A u}{2} - \int_{y \neq 0} (e^{ixu^\top y} - 1 - ixu^\top y \mathbb{1}_{\{\|y\| \leq 1\}}) \nu(dy) \\ &= -ixu^\top b + \frac{x^2 u^\top A u}{2} - \int_{u^\top y \neq 0} (e^{ixu^\top y} - 1 - ixu^\top y \mathbb{1}_{\{\|y\| \leq 1\}}) \nu(dy) \\ &= -ix \left(u^\top b - \int_{u^\top y \neq 0} u^\top y (\mathbb{1}_{\{\|y\| \leq 1\}} - \mathbb{1}_{\{|u^\top y| \leq 1\}}) \nu(dy) \right) \\ &\quad + \frac{x^2 u^\top A u}{2} - \int_{z \neq 0} (e^{ixz} - 1 - ixz \mathbb{1}_{\{|z| \leq 1\}}) \nu \circ \{y \mapsto u^\top y\}^{-1}(dz). \end{aligned}$$

Hence, applying [3, Theorem 3.1] once more shows that $u^\top L$ is a Lévy process with the characteristic exponent κ_u . The converse implication is straightforward by choosing $x = 1$. \square

¹See, e.g., [4, Chapter III, p.124].

Lemma 3.2. *Let $D, D' \in \mathbb{N}$. Assume that W is a D -dimensional Gaussian Lévy process and L is a D' -dimensional purely non-Gaussian Lévy process, both defined on the same probability space. Then W and L are independent.*

Proof. Step 1. We prove that, for any $u \in \mathbb{R}^D$, $v \in \mathbb{R}^{D'}$, two processes $u^\top W$ and $v^\top L$ are independent. Indeed, it is obvious that $u^\top W$ is a Gaussian Lévy process, and $v^\top L$ is a purely non-Gaussian Lévy process due to Lemma 3.1. Denote by $[X, Y]$ the quadratic covariation of two càdlàg real semimartingales X, Y (see, e.g., [4, p.66] or [2, Definition 8.2]). By the bilinearity of quadratic covariation, we get

$$[u^\top W, v^\top L] = \left[\sum_{d=1}^D u^{(d)} W^{(d)}, \sum_{d'=1}^{D'} v^{(d')} L^{(d')} \right] = \sum_{d=1}^D \sum_{d'=1}^{D'} u^{(d)} v^{(d')} [W^{(d)}, L^{(d')}].$$

Since $W^{(d)}$ is continuous and $L^{(d')}$ is purely non-Gaussian, both are Lévy processes null at 0, it implies that $[W^{(d)}, L^{(d')}] = 0$. Hence, $[u^\top W, v^\top L] = 0$. We then apply [2, Theorem 11.43] to get the independence of $u^\top W$ and $v^\top L$ as desired.

Step 2. By choosing a common refinement of partitions, it suffices to prove that $(W_{t_1}, \dots, W_{t_n} - W_{t_{n-1}})$ is independent of $(L_{t_1}, \dots, L_{t_n} - L_{t_{n-1}})$ for all $0 = t_0 < t_1 < \dots < t_n$, $n \in \mathbb{N}$. Let $\{u_k\}_{k=1}^n \subset \mathbb{R}^D$ and $\{v_k\}_{k=1}^n \subset \mathbb{R}^{D'}$ arbitrarily. One has

$$\begin{aligned} I_{(3.1)} &:= \mathbb{E} \left[e^{i \sum_{k=1}^n u_k^\top (W_{t_k} - W_{t_{k-1}}) + i \sum_{k=1}^n v_k^\top (L_{t_k} - L_{t_{k-1}})} \right] \\ &= \mathbb{E} \left[e^{i \sum_{k=1}^n \sum_{d=1}^D u_k^{(d)} (W_{t_k}^{(d)} - W_{t_{k-1}}^{(d)}) + i \sum_{k=1}^n \sum_{d'=1}^{D'} v_k^{(d')} (L_{t_k}^{(d')} - L_{t_{k-1}}^{(d')})} \right]. \end{aligned} \quad (3.1)$$

For $d = 1, \dots, D$, $d' = 1, \dots, D'$ and $k = 1, \dots, n$ we define, for $t \in [0, 1]$,

$$\check{W}_t^{(d,k)} := W_{(t_k - t_{k-1})t + t_{k-1}}^{(d)} - W_{t_{k-1}}^{(d)} \quad \text{and} \quad \check{L}_t^{(d',k)} := L_{(t_k - t_{k-1})t + t_{k-1}}^{(d')} - L_{t_{k-1}}^{(d')},$$

and set $\check{W}_t := (\check{W}_t^{(d,k)})_{1 \leq d \leq D, 1 \leq k \leq n} \in \mathbb{R}^{D \times n}$, $\check{L}_t := (\check{L}_t^{(d',k)})_{1 \leq d' \leq D', 1 \leq k \leq n} \in \mathbb{R}^{D' \times n}$. We now show that $\check{L} = (\check{L}_t)_{t \in [0,1]}$ is an $\mathbb{R}^{D' \times n}$ -valued purely non-Gaussian Lévy process. For any $v \in \mathbb{R}^{D' \times n}$ and $t \in [0, 1]$, one has

$$\text{tr}[v^\top \check{L}_t] = \sum_{k=1}^n \sum_{d'=1}^{D'} v^{(d',k)} \check{L}_t^{(d',k)} = \sum_{k=1}^n \sum_{d'=1}^{D'} v^{(d',k)} (L_{(t_k - t_{k-1})t + t_{k-1}}^{(d')} - L_{t_{k-1}}^{(d')}).$$

For each $k = 1, \dots, n$, since $[0, 1] \ni t \mapsto L_{(t_k - t_{k-1})t + t_{k-1}} - L_{t_{k-1}}$ is a D' -dimensional purely non-Gaussian Lévy process, it follows from Lemma 3.1 that

$$[0, 1] \ni t \mapsto \sum_{d'=1}^{D'} v^{(d',k)} (L_{(t_k - t_{k-1})t + t_{k-1}}^{(d')} - L_{t_{k-1}}^{(d')})$$

is also a purely non-Gaussian Lévy process. Since L has independent increments, we infer that $(\text{tr}[v^\top \check{L}_t])_{t \in [0,1]}$ is again a purely non-Gaussian Lévy process. Analogously, $(\check{W}_t)_{t \in [0,1]}$ is an $\mathbb{R}^{D \times n}$ -valued Gaussian Lévy process without drift. By vectorization and applying Step 1 we get that $(\text{tr}[u^\top \check{W}_t])_{t \in [0,1]}$ is independent of $(\text{tr}[v^\top \check{L}_t])_{t \in [0,1]}$ for any $u \in \mathbb{R}^{D \times n}$ and $v \in \mathbb{R}^{D' \times n}$. Therefore, choosing particularly $t = 1$ yields

$$I_{(3.1)} = \mathbb{E} \left[e^{i \sum_{k=1}^n \sum_{d=1}^D u_k^{(d)} (W_{t_k}^{(d)} - W_{t_{k-1}}^{(d)})} \right] \mathbb{E} \left[e^{i \sum_{k=1}^n \sum_{d'=1}^{D'} v_k^{(d')} (L_{t_k}^{(d')} - L_{t_{k-1}}^{(d')})} \right],$$

which implies the desired conclusion. \square

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